Vikram Sarkar

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The **density** of any subset $A \subseteq \mathbb{N}$ is defined to be

$$\lim_{N\to\infty}\frac{|A\cap\{1,2,\cdots,N\}|}{N},$$

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$$g(n) = O(f(n))$$

if there exists some constant c for which $g(n) \leq cf(n)$ for all $n \in \mathbb{N}$.

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We say $g(n) \ll f(n)$ if g(n) = O(f(n)). For two fixed constants a and b, we say $a \ll b$ instead of a < b to indicate that a is "sufficiently less than b."

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Question

How big is $\omega(n)$? That is, roughly speaking, how large do we expect it to be given, say, n is some fixed order of magnitude? (More rigorously speaking, what is an "average order" of $\omega(n)$?)

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This won't be very central to the main topic, but just to make this notion of "largeness" concrete: f and g have the same **average order** if

$$\lim_{x\to\infty}\frac{\sum\limits_{n\leq x}f(n)}{\sum\limits_{n\leq x}g(n)}=1.$$

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$$\lim_{x \to \infty} \frac{\sum\limits_{n \le x} f(n)}{\sum\limits_{n < x} g(n)} = 1.$$

Answer to the question

The answer is that $\omega(n) \approx \log \log n$. Meaning, they have the same average order.

Here's a heuristic argument to why $\omega(n)$ should be around $\log \log n$.

• For each prime p, let $\chi_p = \begin{cases} 1 & p \mid n \\ 0 & p \nmid n \end{cases}$.

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We can use this logic to prove that $\log \log n$ and $\omega(n)$ have the same average order. (Exercise for the reader.)

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Theorem (Hardy and Ramanujan, 1917)

The **normal order** of $\omega(n)$ is $\log \log n$, that is, for all $\varepsilon > 0$, that is, the density of all $n \in \mathbb{N}$ for which

$$\left| \frac{\omega(n)}{\log \log n} - 1 \right| < \varepsilon$$

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Theorem (Turán, 1934)

$$\sum_{n \le x} (\omega(n) - \log \log n)^2 = (1 + o(1))x \log \log x.$$

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Theorem (Erdős, Kac, 1959)

If a < b are in $\mathbb{R} \cup \{-\infty, \infty\}$, the density of the set

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Okay, what?

Small Simplification

It is not hard to show that it suffices to show the density of the set

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instead. Since $\log \log N = \sum_{p \le N} p^{-1} + O(1)$, it is also not hard to show that if $A(N) = \sum_{p \le N} p^{-1}$, it suffices to show that the density of the set

$$\left\{ n \in \mathbb{N} \mid a \le \frac{\omega(n) - A(N)}{A(N)^{1/2}} \le b \right\}.$$

is $\frac{1}{\sqrt{2\pi}} \int_a^b e^{-x^2/2} dx$ as well.

Definition

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• A graph of φ is shown below.

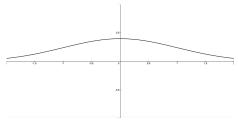
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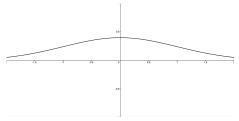
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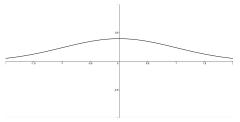
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- Looks like a "squished" bell curve.
- Note that $\int_{-\infty}^{\infty} \varphi(x) dx = 1$ due to the infamous $\int_{-\infty}^{\infty} e^{-x^2} dx = \sqrt{\pi}$.

Definition

A discrete random variable X has a finite sample space S and fixed probabilities of being each element of S. For example, S could be $\{1,2,3\}$, and $\operatorname{Prob}(X=1)=\frac{1}{3}$, $\operatorname{Prob}(X=2)=\frac{1}{6}$, $\operatorname{Prob}(X=3)=\frac{1}{2}$.

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A continuous random variable X which has an infinite sample space $S \subseteq \mathbb{R}$ that is equipped with a **probability density function** f has

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In our situation, we have the sequence of discrete random variables $\{X_N\}$, for which X_N has sample space $\left\{\frac{\omega(n)-A(N)}{A(N)^{1/2}}\mid n\leq N\right\}$, and each $n\leq N$ has equal probability to be chosen.

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$$\mathbb{E}[X] := \frac{1}{|S|} \sum_{s \in S} s \cdot \text{Prob}(X = s).$$

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Theorem (Well-known)

The standard normal distribution is completely determined by all its moments. That is, if there exists some random variable X with

$$\mathbb{E}[X^k] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x^k e^{-x^2/2} \, \mathrm{d}x,$$

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Theorem (Well-known)

Suppose $X_1, X_2, \dots, X_n, \dots$ is a sequence of random variables. Let X be a random variable that is completely determined by all its moments. Then, if $\mathbb{E}[X_n^k] \to \mathbb{E}[X^k]$ for all $k \geq 1$, then

 $\operatorname{Prob}(a \leq X_n \leq b) \to \operatorname{Prob}(a \leq X \leq b) \text{ for all } a < b \in \mathbb{R} \cup \{-\infty, \infty\}.$

• Now, let X_N be the random variable with sample space $\left\{\frac{\omega(n)-A(N)}{A(N)^{1/2}} \mid n \leq N\right\}$ for which each element of this set has an equal probability of being chosen (if there are somehow repeats, then the probability of $X_n = r$ would be the # of n for which $\frac{\omega(n)-A(N)}{\sqrt{A(N)}} = r$ divided by N).

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- Let X be a random variable with probability density function φ . Then, if we had

$$\mathbb{E}[X_N^k] \to \mathbb{E}[X^k] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x^k e^{-x^2/2} \, \mathrm{d}x,$$

we would then have

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• However, note that $\lim_{N\to\infty} (a \leq X_N \leq b)$ is also the density of $\{n \in \mathbb{N} \mid \frac{\omega(n) - A(N)}{\sqrt{A(N)}}\}$. So showing that the moments of X_N approach the moments of X suffices.



$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x^k e^{-x^2/2}$$

Theorem

We have

$$\mu_k := \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x^k e^{-x^2/2} = \begin{cases} (k-1)!! & k \text{ even, } k > 2\\ 0 & k \text{ odd} \\ 1 & k = 0 \end{cases}.$$

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This is not very hard to prove; if k is odd then we are integrating a bounded odd function over a symmetric interval so the integral is 0. If k is even, use integration by parts to get a recurrence between μ_k and μ_{k+2} . So, we want to prove $\mathbb{E}[X_N^k] \to \mu_k$.

Simplifications

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• Note that

$$\mathbb{E}[X_N^k] = \frac{1}{N} \sum_{n \le N} \left(\frac{\omega(n) - A(N)}{A(N)^{1/2}} \right)^k.$$

So if this converges to μ_k , then it equals $\mu_k + o(1)$, where the o(1) is with respect to N, not k.

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So if this converges to μ_k , then it equals $\mu_k + o(1)$, where the o(1) is with respect to N, not k.

• So, we want to show that

$$\frac{1}{N} \sum_{n < N} \left(\frac{\omega(n) - A(N)}{A(N)^{1/2}} \right)^k = \mu_k + o(1).$$

Upon rearrangement, we want to show that

$$\sum_{n \le N} (\omega(n) - A(N))^k = N(A(N))^{k/2} \mu_k (1 + o(1)).$$

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Define
$$L_p(n) := \begin{cases} -1/p + 1 & p \mid n \\ -1/p & p \nmid n \end{cases}$$
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$L_p(n)$

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The important thing here is that $\sum_{p\leq N} L_p(n) = \omega(n) - A(N)$. Therefore, we wish to show that

$$\sum_{n \le N} \left(\sum_{p \le N} L_p(n) \right)^k = N(A(N))^{k/2} \mu_k (1 + o(1)).$$

Now, for each prime p, let L(p) be a random variable for which $L(p)=-\frac{1}{p}$ with probability $\frac{1}{p}$ and $L(p)=1-\frac{1}{p}$ with probability $1-\frac{1}{p}$.

Now, for each prime p, let L(p) be a random variable for which $L(p) = -\frac{1}{p}$ with probability $\frac{1}{p}$ and $L(p) = 1 - \frac{1}{p}$ with probability $1 - \frac{1}{p}$. The following two theorems are not very hard to prove, for the sake of time I won't be going over their proofs.

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Theorem

Suppose $N \ge 1$, and $N_S = N^{1/\log(\sqrt{\log \log N} + 3)}$. Then,

$$\sum_{p \le N} L_p(n) = \sum_{p \le N_S} L_p(n) + O(\log(\sqrt{\log\log N} + 3)),$$

for all n < N.



Relating L_p to L(p)

Theorem

Let $\pi(n)$ be the number of primes $\leq n$. Then, for any positive integer k, we have

$$\sum_{n \le N} \left(\sum_{p \le N_S} L_p(n) \right)^k = N \mathbb{E} \left[\left(\sum_{p \le N_S} L(p) \right)^k \right] + O(3^k \pi(N_S)^k)$$

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So we have expressed something very similar to what we want (the LHS) as something relating the random variable L(p). Right now, the right hand side is not very useful since the main expression is stuck in an expected value. So we need a good way to approximate it.

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Proof

It is clear that M_X is analytic on \mathbb{C} . Note that from the Taylor Series expansion of e^x ,

$$M_X(z) = \mathbb{E}[e^{zX}] = \mathbb{E}\left[\sum_{k \ge 0} \frac{(zX)^k}{k!}\right] = \sum_{k \ge 0} \mathbb{E}\left[\frac{(zX)^k}{k!}\right] = \sum_{k \ge 0} \frac{\mathbb{E}[X^k]}{k!} z^k.$$

Thus $M_X^{(k)}(0) = \mathbb{E}[X^k]$, as desired.

• Now, let $B_N = \sum_{p \leq N} L(p)$. Then, let $M_{B_N}(t) := \mathbb{E}[e^{tB_N}]$ for some $t \in \mathbb{C}$.

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• Now, for $|t| \leq \frac{1}{2}$, we can say by Taylor Series that

$$e^{tL(p)} \! = \! 1 \! + \! tL(p) \! + \! \tfrac{t^2}{2} L(p)^2 \! + \! O\!\left(|t|^3|L(p)|^3\right),$$

SO

$$\mathbb{E}\!\left[e^{tL(p)}\right]\!\!=\!\!1\!+\!\tfrac{p-1}{2p^2}t^2\!+\!O\!\left(\tfrac{|t|^3}{p}\right)\!.$$

• One can then take the product over all $p \leq N$ and then bound to obtain that

$$M_{B_N}\left(\frac{t}{\sqrt{\log\log N}}\right) \to e^{t^2/2}$$

uniformly on some disk centered at 0. Then, by complex analysis magic, the derivatives converge uniformly as well.

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uniformly on some disk centered at 0. Then, by complex analysis magic, the derivatives converge uniformly as well.

• Then we can compute

$$\mathbb{E}[B_N^k] = M_{B_N}^{(k)}(0) = (1 + o(1))\mu_k(\log\log N)^{k/2},$$

since the $k^{\rm th}$ derivative of $e^{x^2/2}$ at 0 is μ_k (!!) from the Taylor Series expansion of e^x .

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Theorem

$$\sum_{n \le N} \left(\sum_{p \le N} L_p(n) \right)^k - \sum_{n \le N} \left(\sum_{p \le N_S} L_p(n) \right)^k = o(N(\log \log N)^{k/2}).$$

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Now, we can use the result we obtained above to get the sum for $p \leq z$.

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The Finish (Continued)

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Note that

$$\begin{split} (\log\log N)^{1/2} - (\log\log N_S)^{1/2} &\leq (\log\log N - \log\log N_S)^{1/2} \\ &= \left(\log\log N - \log\left(\frac{\log N}{\log(\sqrt{\log\log N} + 3)}\right)\right)^{1/2} \\ &= (\log\log(\sqrt{\log\log N} + 3))^{1/2} \end{split}$$

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so by the Binomial Theorem,

$$(\log\log N)^{k/2} - (\log\log N_S)^{k/2} = O\left(k(\log\log N_S)^{\frac{k-1}{2}} \left(\log\log(\sqrt{\log\log N} + 3)\right)\right),$$

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Putting it altogether, we have

$$\begin{split} \sum_{n \leq N} \left(\sum_{p \leq N} L_p(n) \right)^k &= (1 + o(1)) N(\log \log z)^{k/2} \mu_k + O(3^k \pi(N_S)^k) \\ &\quad + o(N(\log \log N)^{k/2}) \\ &= (1 + o(1)) N(\log \log N)^{k/2} \mu_k + O(k(\log \log z)^{\frac{k-1}{2}}) + \cdots \\ &= (1 + o(1)) N(\log \log N)^{k/2} \mu_k \\ &= (1 + o(1)) N(A(N))^{k/2} \mu_k, \end{split}$$

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so we are done. \blacksquare

Thanks!

Any questions?